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Working Memory Tasks Differ in Factor Structure Across Age Cohorts: Implications for Dedifferentiation

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Running head: Working memory used inconsistently with age

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ABSTRACT

Researchers interested in working memory have debated whether it should be considered a single latent cognitive ability or a set of essentially independent latent abilities distinguished by domain-specific memory and/or processing resources. Simultaneously, researchers interested in cognitive aging have established that there are substantial differences in rates of change in various aspects of cognitive function with age. In general, so-called fluid measures of cognitive function including working memory decline at faster rates in later adulthood than so-called crystallized measures. Using an internet working and short-term memory test battery completed by over 95,000 people aged 18-90, we used multiple-group confirmatory factor analysis to assess the extent to which working memory could be considered a single latent ability as well as how its common and unique variance components varied with age. Results indicated a single latent factor, but this factor was not measured consistently across age groups. Both individual test residual variances and factor intercepts showed different patterns of variation with age. We discuss the implications for understanding age differences in working memory function. .

Individuals' performances on tests of different mental abilities are positively inter-correlated. The general factor that characterizes these correlations was observed around the time the first mental ability tests were developed, and has come to be known as Spearman's *g*, after one of the first psychometricians to study it (Spearman, 1904; 1927). Spearman focused his analytical attention on the tests themselves and the extent to which specific tests reflected this general factor (Spearman, 1927), but he also developed hypotheses about the psychological natures of the traits represented by the tests. He proposed that any individual's performance on any cognitive task arose from a general ability (*g*) that could also be applied to any other cognitive task, and a cognitive ability specific to that particular task. Almost immediately, others took issue with this proposition (beginning with Thurstone, 1931), maintaining that mental abilities are distinct and independent from each other. Despite the strong evidence for the pervasiveness of the general factor across test batteries and samples that has developed since then regarding *g* (Jensen, 1998), the debate has continued. This is most notable in studies of individual differences in working memory, which has been shown to correlate highly with a wide range of fluid ability assessments and in particular with '*g*' (e.g. Conway, Kane & Engle, 2003; Cowan, 2005; Kane & Engle, 2002; Unsworth, Redick, Heitz, Broadway, & Engle, 2009).

In the context of working memory, the debate has been fueled by the large literature that shows the separability of a wide range of tasks through experimental manipulations (see reviews in Baddeley, 2007; Conway, Jarrold, Kane, Miyake & Towse, 2007; Logie & van der Meulen, 2009), as well as by studies of individual differences (e.g. Bayliss, Jarrold, Gunn & Baddeley, 2003; Friedman et al., 2008; Miyake, Friedman, Emerson, Witzki & Howerter, 2000). Case studies of the specificity of deficits experienced by victims of neurological damage (reviewed in e.g. Alloway & Gathercole, 2006; Baddeley, Kopelman & Wilson, 2002; Logie & Della Sala, 2005; Vallar & Shallice, 1990) provide additional direct evidence, and the consistently observed fact that different kinds of mental abilities show very different patterns of change with age (e.g., Maylor & Logie, in press; Park et al., 2002; Salthouse, 1995) provides further indirect evidence. Though initially focused on the question of the separability or generality of cognitive ability broadly construed, more narrowly defined aspects of cognition such as executive function, working memory, and attention have been drawn into the separability versus generality debate as well. For example, Baddeley & Logie (1999) posited a multiple-component model of working memory. Under

their model, working memory consists of a central executive controller, one separable subsidiary system for temporary storage of phonologically-based material, and a second separable subsidiary system for temporary storage of visuospatially-based material. In actual task performance at the level of day-to-day activities or even tests of mental ability as usually construed, the central executive coordinates the subsidiary memory systems, controls long-term memory storage and retrieval processes and placement of attention, and carries out the manipulation of the material stored in the subsidiary memory systems. In contrast, other researchers (e.g. Barrouillet, Bernardin, Portrat, Vergauwe & Camos, 2007; Cowan, 2005) have suggested a model based on centrality of processing that emphasizes the interference between verbal and spatial processing that takes place when attentional resources are constrained.

At the same time, the different patterns of cognitive change with age have received considerable research attention. In both cross-sectional and longitudinal data, tests that rely on factual knowledge tend to show gradual average increases until even as late as age 70, and then relatively slow declines with age. In contrast, tests of nonverbal reasoning tend to show faster declines with age that begin around age 30, and tests of perceptual speed show even sharper declines beginning as early as age 20. (Hunt, 1949; Jones & Conrad, 1933; Lovden, Ghisletta, & Lindenberger, 2004; Logie & Maylor, 2009; Salthouse, 2009; Schaie, 1994). In all cases, however, longitudinal data show less extreme differences with age than do cross-sectional data. Whether cross-sectional or longitudinal, the age differences are thought by many to reflect cognitive aging processes, with the differences in rates of change with age reflecting differences between the aging processes underlying tests that measure primarily current efficiency of various kinds of information processing and memory and other tests that measure cumulative products of processing carried out in the past (Salthouse, 2006). To some, these differences suggest the presence of relatively modular mental capacities that change with age at different rates regardless of a common factor such as general working memory capacity (e.g., Logie, Della Sala, MacPherson, & Cooper, 2007; Logie & Maylor, 2009; Park et al., 2002; Perfect & Maylor, 2000). Others, however, see changes in *g* as the main drivers of cognitive aging (e.g., Gow, et al., 2008; McGue & Christensen, 2002; Rabbitt, Lunn, Wong, & Coban, 2008). If changes in *g* are to be understood as the main drivers, *g* and its components such as working memory must be measured consistently with age. To date whether this is the case has not been

investigated systematically throughout the adult lifespan, but it was the primary purpose of the present study.

Consistency of measurement of *g* and its components with age has relevance for the age differentiation hypothesis as well. This is the idea that, during childhood, general ability gradually develops into more specific abilities, while in later life the more specific abilities face global biological constraints that tend to cause their distinctions to blur, that is, later-life dedifferentiation (Balinsky, 1941; Baltes & Lindenberger, 1997; Garrett, 1938; Li et al., 2004; Lovden, Ghisletta, & Sikstrom, 2004). Support for this proposition has been mixed at both ends of the lifespan (e.g., Anstey, Hofer, & Luscey, 2003; Bickley, Keith, & Wolfle, 1995; DeFrias, Lovden, Lindenberger, & Nilsson, 2007; Juan-Espinosa, et al., 2002; Tucker-Drob, 2009; Tucker-Drob & Salthouse, 2009 for studies evaluating later-life dedifferentiation). Variance in measurement properties can provide support for either differentiation or dedifferentiation. Support for the later-life dedifferentiation hypothesis would be provided if there were particular kinds of inconsistencies of measurement of *g* and its components such as working memory in different adult age groups. The inconsistencies should suggest that individual tests become more closely related to each other with age, and/or that individual tests show progressively less test-specific variance with age, especially if this is true to different degrees for different tests.

Establishing Consistent Measurement Across Groups of People

As noted, our main goal in this paper was to evaluate the separability or generality of the working memory construct by assessing how psychometric tests measure the construct across age. To do this, we used confirmatory factor analytic tests of the factor structure of a group of working memory tasks, and of the measurement invariance (Meredith, 1993) of this structure across age groups reflecting the adult lifespan. When a construct is *measurement invariant* across some group division, it is measured in the same way from group to group, and both within-group individual differences and between-group differences in means can be considered to reflect differences in level of the underlying latent construct (Lubke, Dolan, Kelderman, & Mellenburgh, 2003; Meredith, 1993). When a construct is not measurement invariant across groups, between-group mean differences do not reflect differences in level of the underlying construct alone; in addition to construct differences among the groups, there are differences in the relative importance of the various marker variables used to define the latent construct (Hofer, Horn, &

Eber, 1997), and/or some or all of the tests may be somehow measuring different ways of approaching the tasks. When measurement invariance fails, the methodological work involved in addressing its extent and source can be quite informative in identifying to what degree, in what form, and on what tests measurement varies across the groups, leading to both further theoretical developments to improve understanding of mental ability and practical applications to improve tests.

The methodological work in addressing measurement invariance takes place in several steps that impose increasingly restrictive equality constraints on the models used to describe the data in the groups under consideration. Each step is important for establishing full measurement invariance, but the implications of each step for the measurement properties across groups differ. There is a logical order to the steps, and they are carried out in the order given.

The first step is to define a basic model of the factor pattern to describe the data in all groups. This is generally termed establishing *configural invariance*. If this cannot be accomplished, not only is there no point in progressing to the second step, but the basic factor structure of the data is not the same in the groups being compared: there may be more factors in one group than another, or certain items or scales may load on completely different factors in the groups. The second step is to constrain the factor loadings equal across the groups, termed *metric invariance*. When this can be done without loss of overall model fit, the relations among the variables (though not necessarily the factor variances themselves) are the same across the groups and the relative contributions of the various tests to the factor(s) are the same in all groups. When it cannot be done, there are important differences in the extent to which some or all of the tests represent the factors. Next, residual variances are constrained equal. When this is possible, the reliabilities of the tests are functions only of the factor variances in all groups. When it is not possible, there are differences either in error variance or in systematic test-specific variance across groups. Alternatively, residual variances may be left free and factor intercepts constrained equal to test *strong invariance*. When this is possible, latent factor mean differences are interpretable, though mean differences in the tests themselves may not be attributable to mean differences in the latent factor(s). Finally, factor intercepts and residual variances are constrained equal, and *strict invariance* is tested. When this can be done without loss of overall model fit, group mean differences in the tests can be attributed to mean differences in the latent factor(s), and the tests can be considered to reflect the latent

construct in the same way in all groups. When there are differences in intercepts but not residual variances among groups, there is some difference between groups that affects mean test levels but not the latent factor(s) (Brand, 1987), such as different problem-solving strategies, levels of background knowledge or problem-solving sophistication, or familiarity with procedures (Johnson & Bouchard, 2007b; Wicherts, 2007; Wicherts, Dolan, & Hessen, 2005). The tests involved in establishing measurement invariance tend to have low power (Molenaar, Dolan, & Wicherts, in press). That is, there may be important group differences in measurement properties that cannot be detected with these statistical tests. Thus, working with large samples is important.

The Present Study

The purpose of the present study was to investigate systematically the extent to which one component of g , working memory, is a unitary construct measured consistently across the adult lifespan, in order to evaluate the extent to which it may differentiate with age. Age is a particularly important variable on which to evaluate the extent to which working memory represents a single unitary construct because individuals pass through different ages in the course of their lives. This is not true of many grouping variables: for example, in general individuals are either male or female; they are not male at some times and female at others. In addition, any failures of measurement invariance of working memory with age that we observed would provide new evidence for the degree to which the dedifferentiation hypothesis holds in later life. To accomplish these two goals, we made use of multigroup confirmatory factor analysis testing measurement invariance in a battery of tasks put together to address different aspects of working and short-term memory function and administered over the internet in collaboration with the British Broadcasting Corporation (BBC). Ideally, our sample would be longitudinal, so that we could evaluate the extent to which the common working memory factor can be measured consistently in the same people as they change with age. This was not the case for this sample, but the extent to which working memory can be measured consistently in people born at different times has not been evaluated either, so our cross-sectional sample served as a good introduction to the general subject of its consistency of measurement across age.

Our internet collaboration with the BBC made it possible to accumulate an unusually large sample in excess of 95,000 individuals ranging in age from 18 to 90. Collecting psychological data over the

internet is becoming increasingly common (see Birnbaum, 2004; Reips, 2002; Skitka & Sargis, 2006). It makes common the accumulation of large samples, and these samples tend to represent broader demographics than most laboratory-based studies and even than many studies that rely on community volunteers, with substantial savings of research time, energy, and resources. In addition, there may be advantages particular to research on aging. Participants do not have to travel and may be less anxious when tested in their own familiar environments (Maylor & Reimers, 2007), and older adults are increasingly being encouraged to use the internet (Cutler, Hendricks, & Guyer, 2003; Selwyn, Gorard, Furlong, & Madden, 2003). Moreover, accumulating evidence suggests that internet studies can reliably replicate more conventional data collection methods even in studies of aging (Della Sala, Darling & Logie, in press; Logie & Maylor, 2009; Maylor & Logie, in press; Maylor, Reimers, Choi, Collaer, Peters, & Silverman, 2007; Robins, Trzesniewski, Gosling, & Potter, 2002).

METHOD

Sample

Participants were volunteers who accessed the Science pages of the BBC's official website between May 25, 2006 and March 19, 2007. The study was clearly advertised on the site, and it was also advertised on BBC radio programs, was the topic of an article in the August, 2006, *Radio Times* program guide magazine (Logie, 2006), and was featured in a major BBC television documentary on human memory aired on August 9, 2006 (Cadman, 2006). The first participant completed the tests within an hour of their posting on the site and approximately one third of the participants had completed the tests prior to the August television broadcast, suggesting that many participants regularly and spontaneously accessed the BBC site. Defining a data record as the data set for a single attempt to undertake the test battery including partially completed attempts, 160,405 data records were collected during the 10-month data collection period. Participants were requested to provide information on country of residence, highest level of completed education, sex, and age.

Participants reported 156 different countries of residence, but the vast majority reported residing in the United Kingdom or United States. Initial data analysis revealed no differences between those reporting residence in English language-dominated countries and those not. Though participants were not specifically asked to indicate fluency in English, we considered it reasonable to assume that they had a

high level of English fluency because they would have had to find and choose to access the website through English language web pages maintained by the BBC.

For this study, we made use of a subset of the 160,405 data records collected. We excluded all those who did not provide age and/or education, as well as those who reported ages under 18 and over 90. This left 111,497 data records. Among the data records, it was not possible to determine whether individuals completed the tests multiple times. It was, however, possible to ascertain how many times a particular computer had been used to attempt the tests. To minimise the possibility of multiple attempts by the same individual, we made use of only the first data record from each computer for all those 65 and under. There were relatively few participants above age 65, and inspection of the multiple data records from the same computers indicated that few if any were from the same individuals (age, education, and sex did not match and scores varied widely). We therefore included these data records in order to maintain sample sizes as large as possible for the older groups. This left a total of 95,201 data records. Among them were participants who reported unrealistic levels of education such as a postgraduate degree at age 18. We deleted the education variable for these data records. Some participants did not report sex, but we did not exclude their data records on this basis.

Measures

We made use of the five tasks in the data set that had scale scores that could reasonably be considered continuous and measured in scales with roughly equal numbers of intervals. There were several other dichotomous measures in the set (see Logie & Maylor, 2009; Maylor & Logie, in press). Our only reason for not including them in our analyses was that dichotomous variables have very limited measurement properties that would not contribute meaningfully to estimation of a common factor in combination with the other quantitative variables. The five tasks tapped a wide range of aspects of working, visual, and verbal short-term memory. Three of the tests (Digit Span, Working Memory Span and Spatial Orientation) have been used in multiple previously published assessments of fluid intelligence, a fourth (visual pattern span) has been published with normative data. One of the tests (Feature Binding) was more novel, but variations of this test recently have been used to assess cognitive decline with age. Specific descriptions of our variables follow.

Education. Participants reported highest level of attained education on a scale from 1 to 7, where 1 indicated no education; 2 primary education; 3 secondary education; 4 technical or vocational college; 5 other college; 6 college graduate with a first degree; and 7 postgraduate education.

Feature Binding (adapted from Brockmole, Parra, Della Sala & Logie, 2008; Logie, Brockmole & Vandenberg, 2009). Participants were presented with screens containing 1 to 4 objects for 1 second each. The test started with two trials, each consisting of one object, and followed with two trials consisting of two objects, two trials consisting of three objects, and finally two trials consisting of four objects, for a total of 20 objects across all trials. After the presentation of each screen, there was a 1-second gap followed by a test screen showing a selection of colors, shapes, and locations. The task was to report the color, shape, and location of each object that had been on the previously presented screen by clicking on the color and shape of each object and then clicking on its presented location. The test stopped when participants failed accurately to report anything from the two trials at a particular array size.¹ Performance was scored as the number of trials for which color, shape, and location of all objects were reported correctly and thus ranged from 0 to 8. We presumed that specific abilities tested by this task involved visual working memory, visual attention, and binding of features to form representations of integrated objects (e.g. Brockmole et al., 2008; Logie et al., 2009; Mitchell, Johnson, D'Esposito, Raye & Mather, 2000; Treisman, 2006).

Visual Pattern Span (adapted from Logie & Pearson, 1997). Participants were presented with screens containing matrices consisting of patterns of white and blue squares for 2 seconds each. Immediately after presentation of each screen, a matrix of blank squares was presented, and participants were to click on the squares in the matrix that had been blue on the previous screen. There were two trials each of 3x3 (5 blue squares), 3x4 (5 blue squares), 4x4 (8 blue squares), 4x5 (9 blue squares), and 5x5 (9 blue squares), for a total of 10 trials. The test stopped when participants failed to recall all of the blue squares at a single matrix size. Performance was scored as the number of trials for which all blue squares were recalled and thus ranged from 0 to 10. A standard laboratory version of this test has been published with normative data and has been shown to be a robust measure of immediate visual memory (Della Sala, Gray, Baddeley, Allamano & Wilson, 1999).

Digit Span. Participants were presented with sequences of screens containing randomly generated single digits for 1 second each. At the end of each sequence, a screen containing a blank box appeared, and participants were to type in the sequence of digits they had seen. Two sequences at each length from 3 to 9 were presented. The test stopped when participants failed to recall both sequences at a single sequence length. Performance was scored as the total number of digits recalled in the correct orders and thus ranged from 0 to 84. Variations of this test, originally devised by Jacobs (1887), have been included in standard test batteries of fluid intelligence almost since their inception, and appear in contemporary standard assessments such as the Wechsler Adult Intelligence Scale. It is assumed to assess the specific ability to retain a verbal sequence in immediate memory, and is often associated with the operation of the phonological loop component of working memory (see review in Baddeley, 2007).

Working Memory Span (adapted from Baddeley, Logie, Nimmo-Smith and Brereton, 1985; Duff & Logie, 2001). Participants were presented with sequences of screens containing simple sentences such as 'Flies are insects' or 'Mobile phones are made of cheese,' and buttons indicating 'true' or 'false.' As fast as possible, participants were to click on the appropriate button for the presented sentence and were to remember the last word of each sentence. A new sentence was then presented until the sequence was complete. Participants were then presented with a 4x5 array of words in the left two-thirds of the screen. The final words of each sentence in the sequence were randomly distributed throughout the array, and the remaining words were unrelated foils. Participants were to click on the final words from the presented sentences in the correct order and drag them to boxes arranged vertically on the right third of the screen. Two sequences at each length from 2 to 6 were presented. The test stopped if the participant failed to recall all of the last words of the sentences at a single sequence length. Performance was scored as the total number of final sentence words recalled in the correct order and thus ranged from 0 to 40. The Baddeley et al. (1985) version of this test was shown to correlate with a range of mental abilities. Different versions of this kind of test, thought to involve both mental processing and immediate memory, have been shown to correlate strongly with fluid intelligence (e.g. Baddeley et al., 1985; Kane & Engle, 2002).

Spatial Orientation (adapted from Logie & Baddeley, 1983). Participants were presented with a series of screens showing male figures facing away or towards the observer and either upright or inverted. Below each figure were buttons indicating 'left' and 'right.'. Each figure had a blue ball in one hand and a

white ball in the other. Participants were to click on the button indicating the hand holding the blue ball. Performance was scored as the number correct in 30 seconds and ranged from 0 to 41. This task is thought to reflect speed of response as well as spatial ability. Originally devised by Benson & Gedye (1963), versions of this task have been used in a wide range of psychometric test batteries, and it has been shown to correlate with fluid intelligence as well as with other measures of speed of processing (e.g. Carter & West, 1985; Turnage & Kennedy, 1992).

Statistical Analyses

The Working Memory Span variable was negatively skewed (-1.033); all others were reasonably normally distributed. We reduced the degree of skew of the Working Memory Span variable by squaring it and then standardized all variables in the full sample, in order to place them on the same scale. We then created 14 age groups beginning with ages 18-20 and extending in 5-year increments (21-25, 26-30, etc.) to 81 and over. In the full sample, the covariance matrix had a single eigenvalue of 2.09 greater than 1; the next largest eigenvalue was .85, and this pattern was very consistent throughout the age groups.² There was no meaningful evidence for anything other than a single general factor that accounted for around 42% of the variance, in the sample as a whole and within each age group. We therefore fit a single-factor model separately to the data for each age group using confirmatory factor analysis as implemented in Mplus (Muthen & Muthen, 1998-2006) and proceeded from there to apply the constraints described above that were needed to test measurement invariance (Meredith, 1993).

To assess relative model fit, we report chi-squared and $-2 \times \log\text{likelihood}$, Comparative Fit Index (CFI), Tucker-Lewis Index (TLI), Root Mean Square Error of Approximation (RMSEA), Akaike Information Criterion (AIC; Akaike, 1983), and Sample-Size Adjusted Bayesian Information Criterion (BIC; Schwartz, 1978). Because the progressively restrictive tests of measurement invariance create nested models, the differences in chi-square and $-2 \times \log\text{likelihood}$ between pairs of models are distributed as chi-square with the differences in degrees of freedom or numbers of estimated parameters as degrees of freedom. In large samples such as ours, however, these tests are generally too restrictive. That is, they show statistically significant differences in model fit due simply to sample size. AIC and BIC are information-theoretic fit statistics that are not subject to this problem. Moreover, they explicitly recognize model parsimony, particularly BIC, and they tend to be more sensitive than CFI, TLI, and RMSEA. In our

application, this means that they tend to indicate preference for models allowing the constraints that indicate measurement invariance, so they provided the strongest tests of failure of measurement invariance and we relied on them most heavily. For larger samples, BIC generally provides better estimates than AIC (Markon & Krueger, 2004). Lower (smaller) fit statistics indicate preferred models for both AIC and BIC, with differences of 10 or more considered substantive for BIC.

RESULTS

Descriptive Statistics

Table 1 and Figure 1 show descriptive statistics for the study variables for each age group. Overall, 60% of participants were female. The female percentage ranged from a low of 42% in the 76-80 age group (the only age group without a female majority) to 65% in the 51-55 age group. The participants tended to be rather young (mean age was 33.55, 12.48 SD) and well educated (mean level of education was 4.97, 1.48 SD, indicating an average close to college educated). The largest age group was 26-30, and participation fell for each age group beyond that. Consistent with the years required to complete the highest levels of education, average education rose in the age groups until the 26-30 age group. Consistent with general demographic trends, average education fell with age after that. Consistent with many other studies, there were also substantial differences in mean scores on the tests across age groups. These were the subjects of the more extensive analyses that followed.

Tests of Measurement Invariance of Working Memory

Table 2 shows the step-by-step results of the tests of measurement invariance of the working memory construct across age groups. The simple one-factor model provided a reasonable description for each age group, establishing configural invariance. Given that BIC was lower (smaller) in the model with factor loadings constrained equal across groups and the other fit statistics gave indications consistent with this, it was also reasonable to consider that factor loadings could be constrained equal across groups, establishing metric invariance. All the fit statistics showed considerable deterioration when we constrained residual variances equal, however, and there was further marked deterioration when we constrained intercepts equal. Thus it was not possible to attain strict measurement invariance across age groups.

There were differences in sex and level of education across the age groups. Any failures of measurement invariance of the working memory construct across education and/or sex could contribute to

the failures of measurement invariance with age because there were small correlations between education and test scores (.09 on average across the 5 tests) and between sex and test scores (-.05 on average across the 5 tests, with females coded as 1, males coded as 0). We therefore also tested invariance of measurement of the working memory construct across levels of education and sex. These results are shown in Table 2 as well. For level of education, it was possible to constrain both factor loadings and residual variances equal across groups, establishing metric invariance, but strict measurement invariance could not be established because intercepts could not be constrained equal. Results across sex were analogous to those for age groups: it was possible to constrain factors loadings equal, but not to constrain either residual variances or intercepts equal in males and females. Differences in levels of education and proportions of females across the age groups may thus have contributed to the failures of measurement invariance of a general working memory factor across age groups.

Evaluating the Sources of Measurement Variance of Working Memory

Though we did not observe strict measurement invariance of working memory across age groups, it was possible to obtain strict measurement invariance across many different combinations of pairs of age groups, especially between those adjacent to each other, such as ages 20 and under and ages 21-25. It was not, however, possible to constrain either intercepts or residuals equal without loss of model fit over what might be considered the peak adult age range of 18-35. We thus thought it of greater interest to examine the patterns of freely estimated residual variances and intercepts across the age groups in order to develop ideas about how general working memory capacity and more specific abilities might be used at different ages to carry out tasks that were presumably reasonably novel to the participants. Table 3 shows these results.

General patterns were clearly evident. We begin with residual variances, reflecting the extent to which specific aspects of working memory and/or measurement error independent of the general construct accounted for individual differences in task performance. For Feature Binding, the residual variances generally rose gradually with age until ages 46-50 when the residual variance was .66, and then began a somewhat more rapid decline, reaching .44 by ages 81 and over. Visual Pattern Span showed the opposite pattern: its residual variance decreased gradually from .69 at ages 18-20 to .44 at ages 66-70, and then increased. Residual variances for Digit Span increased rather steadily from .71 to 1.01. Working

Memory Span residual variances increased from .68 at the youngest ages to .80 in middle age, and then slowly decreased, while residual variances for Spatial Orientation decreased gradually from .92 at ages 20 and under to about .55 in old age. Factor intercepts, reflecting mean levels specific to the individual tests, also showed clear general patterns. Feature Binding intercepts showed continual decreases with age, as did those of Visual Pattern Span and Working Memory Span, though the slopes were obviously different. In contrast, Digit Span and Spatial Orientation intercepts increased until around age 30 before beginning continual decreases throughout the remainder of the lifespan.

All of the patterns of change in both residual variances and intercepts were clear enough that it made sense to fit regression lines to them. Because of the marked differences in sample size, we used regression equations weighted by age group n , thus giving greater weight to the younger age groups with the largest numbers of participants. We considered linear, quadratic, and cubic regression equations. In most cases, quadratic equations fit best. All regressions were highly significant. For residual variances, we used quadratic equations for all but Visual Pattern Span, for which a linear equation was sufficient, accounting for 90% of the variance. For Digit Span, Working Memory Span, and Spatial Orientation, the quadratic regression equations accounted for 63%, 81%, and 98% of the variances, respectively. The n -weighted quadratic regression equation for Feature Binding accounted for only 42% of the variance, but an unweighted quadratic regression equation accounted for 88%. Figure 2 shows the fitted weighted regression lines for the residual variances of the tests. For intercepts, we again used mostly quadratic equations. A linear equation was sufficient for Visual Pattern Span intercepts, accounting for 99% of the variance. Spatial Orientation intercepts required a cubic equation, but it accounted for 98% of the variance. The quadratic equations for Feature Binding, Digit Span, and Working Memory Span accounted for 99%, 86%, and 99% of the variance, respectively. Figure 3 shows the fitted weighted regression lines for the tests' intercepts.

Table 4 shows the intercepts for the education groups, and the residual variances and intercepts for males and females. It would also have been possible to establish measurement invariance across some combinations of educational groups, but again we believed that it was more informative to show the freely estimated intercepts. For all tests, the lowest educational groups showed markedly poorer average performance than the higher educational groups, though sample sizes were much smaller. The highest

overall performance was in college graduates with first degrees: their intercepts were highest for all the tests except Digit Span. Residual variances were generally higher in males than in females, as were intercepts, with the exception of the intercept for Working Memory Span.

DISCUSSION

Our goals in this study were systematically to investigate the extent to which the working memory construct can be considered unitary, and to assess whether its structure can be measured consistently across the adult lifespan. This is important in understanding the extent to which age differences in working memory performance reflect differences in a general working memory capacity or differences in relatively modular working memory components that can function relatively independently. It is also important in understanding the ways in which the relative generality or modularity of working memory may change with age; that is, whether working memory may become less differentiated in later adulthood. Our results indicated, first, that a single latent general factor described the data for each age group. There was no evidence that the various aspects of working memory could be considered truly modular even when taking each age group on its own. At the same time, on the order of 75% of the variance in each individual working memory task was independent of the variance in the other tasks. While perhaps a third of this independent variance might be considered outright error variance, most of the tests used here have been shown previously within specific age groups to be robust and reliable measures of the mental abilities that they purport to measure, so the large amount of test-specific variance indicated that it would be difficult to consider the working memory construct to be truly unitary as well.

The single-factor structure of our tests of working memory could be considered invariant with age. That is, the individual tests were equivalently reliable indicators of that single factor in all age groups. There were, however, more subtle violations of measurement invariance that undermined the consistency of measurement with age, and in particular the ability to consider mean differences in test scores with age to be attributable to overall working memory. The existence of differences with age in residual variances indicated that there may be important differences with age in the extent to which people must rely on a general working memory capacity rather than on specific, relatively modular abilities or skills. For some tasks, there was progressively greater reliance with age on the general capacity, while for other tasks there was progressively greater reliance on task-specific abilities. Just as importantly for our primary

question of interest, these differences could also indicate differences with age in the abilities of the tests to measure their constructs of interest, including the general working memory factor (i.e., reliability). The existence of differences with age in intercepts indicated that age was associated with some difference(s) that affected mean individual task performance levels but not level of general working memory capacity. Possible sources of these differences include development of memory strategies (e.g. Logie, Della Sala, Laiacona, Chalmers & Wynn, 1996), levels of background knowledge or problem-solving sophistication, and familiarity with task procedures.

Possible Objections

Despite the clear advantage of its very large sample size, our study might be subject to two potential objections that should be acknowledged and evaluated before discussing its results in greater detail. The most obvious is the cross-sectional nature of our study's design. We discuss many of our results as if the attribution of effects to age is clear, but of course age and cohort effects (including the Flynn Effect; Flynn, 1994; 2007) were confounded in these data. Systematic longitudinal studies over the age range reported here and with a consistent battery of tests are impractical, but over shorter time periods may be preferred when linking mean differences to aging, but even those studies can understate age effects because of the existence of test practice effects (Rabbitt, Lunn, Wong, & Cobain, 2008; Salthouse, 2009). If the results of longitudinal studies suggesting smaller and later cognitive decline with age than cross-sectional studies (Schaie, 1994) are correct, it may be possible that longitudinal samples would show invariance of measurement with age. In rebuttal, we note that analyses (manuscript in preparation) of three waves of data in the Lothian Birth Cohort 1921 (Deary, Whiteman, Whalley, Fox, & Starr, 2004) showed variance in measurement of general intelligence with age, particularly for the intercepts. Though it is conceivable that this occurred purely because of differential practice effects, it seems unlikely. The Lothian Birth Cohort data spanned the age range from 79 to 87, a much narrower range than in our study but completely contained in our oldest age group. The failure of measurement invariance there is a good indication that measurement invariance would also fail over a broader longitudinal age range, though it is possible that it would not fail for younger longitudinal age range. It is also possible that measurement invariance could fail for some groups of cognitive tasks but not for others.

Second, our results indicated failure of measurement invariance across educational groups as well as age groups, indicating that some of the variance in measurement with age may be attributable to differences in education, not just to differences in educational credentials attained, but also to cohort differences in educational curricula. But some of the measurement variance may also be due to selection in the sample for general intelligence itself that increased with age; that is, participants chose to participate in a computer-administered cognitive test battery and the extent to which they were self-selected for above average age-adjusted general intelligence may have varied (most likely increased) with age. Similar comments can be made with respect to sex differences with age in participation, and the possibility of differences in socioeconomic status and other demographic factors that could impact internet access and experience with computers. Evidence against these alternative accounts however, comes from the observation that there were variances of measurement between just, for example, the 26-30 and 31-35 age groups, for whom cohort differences of all kinds are less credible. Most importantly, whereas such cohort differences limit our ability to draw clear inferences about the age effects, they do not limit our ability to assess the failures of measurement invariance. All of these potential additional reasons for sampling differences that could have created the failures of measurement invariance would also be variables across which measurement of working memory capacity ideally should be invariant, and so these possible sampling differences do not undermine our conclusions in this regard.

Patterns of Residual Variances with Age

There were clear patterns of residual variance with age that have important implications for the theoretical conceptions of general and specific working memory abilities. Feature Binding showed some increases in residual variance until about age 40, followed by decreases after that. Mean performance on this test declined across the full age range, but increasingly steeply after age 40 (see also Brockmole et. al, 2008; Brown & Brockmole, in press). One interpretation of this is that people were able to use test-specific abilities and skills to offset underlying decline in or ineffectiveness of general working memory capacity or the central executive for task purposes until about age 40, but that after that they relied increasingly but ineffectively on general working memory capacity. Table 5 shows the proportions of variance attributable to general working memory capacity for each test in each age group. It is unlikely that such a pattern would result from changes in test reliability, as the pattern would imply that test

reliability increased as performance decreased. Similarly, Visual Pattern Span showed universal decreases in both residual variance and performance, possibly indicating that participants increasingly and ineffectively relied on general working memory capacity for task performance with age. Again, such a pattern would be unlikely to result from changes in test reliability. Digit Span showed universal increases in residual variance with age, and performance increased the longest of any of the tests, until the mid-40's. This suggested that participants were able to make successful use of test-specific abilities and skills to increase or maintain performance at least until middle age. After that, however, participants were apparently less likely to have appropriate test-specific skills and abilities. Increases in residual variance with age would be more likely to be associated with decreasing test reliability, but there is little reason to suspect changes with age in reliability of Digit Span when there is no indication of such changes for the other tests. Working Memory Span showed a pattern very similar to that of Feature Binding, but it extended further, until about age 55. The patterns for Digit Span and for Working Memory Span pattern were consistent with the results of many studies showing better preservation of abilities related to vocabulary (e.g., Cattell, 1971; Horn, 1986). Spatial Orientation showed a pattern very similar to that of Visual Pattern Span, except that presumed increasing reliance on general working memory capacity appeared to have some beneficial effect on performance at least until about age 30 for Spatial Orientation. Thus it is clear that variance in performance on the tests in our battery was not consistently attributable to general or specific memory capacities across the age groups.

There was a marked difference between the patterns of residual variance for tests that differentially rely on verbal and visuospatial abilities. Residuals for Feature Binding, Visual Pattern Span, and Spatial Orientation showed decreases across the later lifespan whereas those for Digit Span and Working Memory Span either remained relatively stable or increased. Consistent with many other studies, the relative similarities in the patterns with age among the verbally oriented tasks and among the visuospatially oriented tasks, along with the differences between these two sets of patterns, suggest a basic verbal-visuospatial division that is consistent with working memory comprising multiple, domain-specific resources (e.g., Baddeley, 2007; Baddeley & Logie, 1999; Johnson & Bouchard, 2007a,b; Logie & van der Meulen, 2009; Saito, Logie, Morita & Law, 2008), that also show different age-related trajectories (Logie & Maylor, 2009) rather than a domain-general system (e.g. Barrouillet et al., Cowan, 2005). A

possible caveat is that these differential age trajectories could reflect cohort differences in experience with verbal and visuospatial kinds of problems such as would be the case if educational curricula and recreational activities have increasingly emphasized visuospatial tasks over the last 50 years or so. This latter explanation is, however, less convincing as an account of the very different patterns of change for Digit Span and Visual Pattern Span between the large participant groups that are relatively close in age and would have had very similar educational and recreational experiences such as between 21-25 and 26-30, or between 31-35 and 36-40.

Overall, the patterns of decreasing residual variance with age for most of the tests tended to support the dedifferentiation hypothesis in adulthood. Stronger support for this hypothesis, however, would have been provided by increases in test loadings on the general working memory factor with age. Moreover, in the very oldest age groups there was some evidence for increases in residual variances, contradicting the dedifferentiation hypothesis. These indications should be considered tentative because the sample sizes were smaller in those age groups than in the other groups. Nevertheless, the numbers in those older groups ($n=1698$ for age > 65 years) were larger than in many previous studies of cognitive changes with age.

Patterns of Differences in Intercepts with Age

Digit Span and Spatial Orientation were the two tests that showed increases in performance at least until about age 30. Performance on Spatial Orientation was also the best maintained even once decline with age began. This might reflect the fact that this was the task that had the greatest probability of generating correct answers completely by chance, as participants had only to choose between left and right. Working Memory Span probably relied most on verbal knowledge and, after Digit Span and Spatial Orientation, performance on it was best maintained with age. In contrast, tasks that rely more on visual processing, namely Visual Pattern Span and Feature Binding, showed the sharpest decreases in performance with age. These might have been the tests with which participants overall were least likely to have some experience and the most likely to show cohort differences in experience. However, again this is not particularly convincing as the sole explanation because there was measurement variance between large groups adjacent in age, such as 26-30 and 31-35, and 45-50 and 51-55, for which cohort differences were less likely. The more dramatic declines in visual working memory tasks in this large sample lend

support to lab-based studies suggesting that age-related working memory deficits for visuospatial material are more severe than those observed for verbal material (Jenkins, Myerson, Hale, & Fry, 1999; Jenkins, Myerson, Joerding, & Hale, 2000; Leonards, Ibanez, & Giannakopoulos, 2002; Myerson, Hale, Rhee, & Jenkins, 1999), and they counter observations to the contrary (Park, Lautenschlager, Hedden, Davidson, Smith, & Smith, 2002; Salthouse, 1995). The results also are consistent with the suggestion that there is greater rather than lesser differentiation among cognitive abilities with age (e.g. Park et al., 2002).

Conclusions: Implications for Understanding Working Memory and its Specific Components and for Future Research

At all ages in our data, a single factor operated strongly across all five working memory tasks, indicating that some form of general capacity or central executive contributed to all of them, though to varying degrees. Moreover, the extent to which the five tasks either contributed to or were represented by this general factor was consistent at all ages. Though they should be replicated in other studies subject to different limitations, the failures of measurement invariance across age in this study tended to undermine rather than support the dedifferentiation hypothesis regarding cognitive changes with age: the general working memory factor did not appear to contribute consistently to test performance with age, as evidenced by the differences in residual variances and intercepts with age. These differences suggest that, not only do relatively specific abilities change in different ways with age, but people make use of their general and specific working memory capacities differently with age or developmental experiences. Exploration of these themes in future research could help to develop both more effective educational techniques and targeted interventions to help older adults cope with tasks that are most likely to be affected by cognitive decline.

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FOOTNOTES

¹This was purely for practical reasons, to prevent participant frustration. As each step of the tasks increased in difficulty, it should have had no impact on participants' scores, as participants who could not recall all the features of two objects in either of two trials presumably would also not have been able to recall all the features of three trials. Participants were scored for all the trials for which they gave correct answers, regardless of how far within the full test they progressed. This was true for Visual Pattern Span as well.

²Eigenvalues, or characteristic roots, of a matrix summarize the extent to which the variance in the matrix can be consolidated or reduced to underlying dimensions, the primary purpose of factor analysis. A 5x5 variance-covariance matrix such as we used here generates five eigenvalues that might commonly range in size from 3 to .2, though the presence of no single eigenvalue as high as 2 would not be uncommon. A standard rule of thumb in factor analysis is that the data contain as many factors as there are eigenvalues greater than 1. Many studies have documented that this rule of thumb is biased toward the extraction of too many factors because eigenvalues often cluster just above 1, and alternative methods have been developed to produce unbiased estimates of the appropriate number of factors. As the rule of thumb method provided clear evidence of the existence of a single factor in our data and there was no room for upward bias in this estimate (the only lower alternative was 0), there was no point in applying any of these other methods to determine the appropriate number of factors.

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Table 1
Descriptive Statistics of Study Variables

	<u>Age</u>	<u>Education</u>	<u>Feature Binding</u>	<u>Visual Pattern Span</u>	<u>Digit Span</u>	<u>Working Memory Span</u>	<u>Spatial Orientation</u>
<u>1. Ages 18-20 (N=11,595, 64% Female)</u>							
Mean	18.93	3.72	4.71	5.53	53.95	27.48	8.64
SD	.83	.95	1.77	1.57	18.08	9.11	3.51
Skewness	.13	.34	-.53	-.02	-.15	-1.30	.03
<u>2. Ages 21-25 (N=19,228, 60% Female)</u>							
Mean	23.05	5.03	4.61	5.42	56.42	27.28	8.81
SD	1.41	1.31	1.84	1.56	18.29	9.22	3.53
Skewness	-.05	-.50	-.50	.01	-.34	-1.23	-.01
<u>3. Ages 26-30 (N=17,146, 58% Female)</u>							
Mean	27.88	5.47	4.46	5.20	57.00	26.92	8.92
SD	1.43	1.36	1.84	1.56	18.24	9.39	3.44
Skewness	.13	-.75	-.47	.06	-.39	-1.16	-.08
<u>4. Ages 31-35 (N=12,837, 56% Female)</u>							
Mean	32.95	5.33	4.33	5.00	57.32	26.43	8.87
SD	1.41	1.46	1.85	1.52	18.32	9.60	3.36
Skewness	.05	-.53	-.43	.07	-.43	-1.10	-.11
<u>5. Ages 36-40 (N=9,772, 56% Female)</u>							
Mean	37.88	5.09	4.16	4.75	56.52	26.04	8.68
SD	1.42	1.52	1.85	1.52	18.79	9.72	3.25
Skewness	.12	-.26	-.32	.06	-.45	-1.00	-.13
<u>6. Ages 41-45 (N=7,186, 59% Female)</u>							
Mean	42.93	5.00	3.89	4.45	56.02	25.38	8.33
SD	1.41	1.54	1.85	1.47	18.53	10.06	3.10
Skewness	.08	-.14	-.23	.13	-.38	-.88	-.10
<u>7. Ages 46-50 (N=6,094, 63% Female)</u>							
Mean	47.96	4.92	3.64	4.15	55.33	224.57	8.09
SD	1.42	1.53	1.87	1.48	18.57	10.34	3.01
Skewness	.05	-.07	-.14	.10	-.38	-.78	-.06
<u>8. Ages 51-55 (N=4,737, 65% Female)</u>							
Mean	52.88	4.90	3.37	3.91	54.54	23.94	7.95
SD	1.40	1.54	1.82	1.42	18.82	10.45	2.95
Skewness	.13	-.04	-.05	.08	-.33	-.67	-.01
<u>9. Ages 56-60 (N=3,166, 63% Female)</u>							
Mean	57.88	4.82	3.14	3.67	54.18	23.57	7.62
SD	1.38	1.57	1.85	1.39	18.78	10.54	2.88
Skewness	.11	.00	.09	.02	-.33	-.63	-.06
<u>10. Ages 61-65 (N=1,492, 62% Female)</u>							
Mean	62.72	4.74	2.79	3.42	52.43	22.17	7.18
SD	1.39	1.60	1.79	1.43	19.25	10.83	2.80
Skewness	.26	.09	.12	.17	-.27	-.52	.07
<u>11. Ages 66-70 (N=947, 59% Female)</u>							
Mean	67.82	4.46	2.34	3.04	48.72	19.71	6.52
SD	1.40	1.58	1.71	1.34	19.93	11.12	2.79

Skewness	.17	.29	.29	.02	-.23	-.24	.21
<u>12. Ages 71-75 (N=447, 52% Female)</u>							
Mean	72.87	4.55	2.08	2.91	46.06	17.46	6.32
SD	1.40	1.59	1.66	1.37	18.73	11.05	2.68
Skewness	.14	.18	.61	.04	-.03	.04	.04
<u>13. Ages 76-80 (N=210, 42% Female)</u>							
Mean	77.85	4.30	2.10	2.75	46.95	17.18	6.14
SD	1.45	1.75	1.69	1.58	17.49	11.58	2.60
Skewness	.19	.15	.51	.89	-.06	.08	-.04
N							
<u>14. Ages 81 and Over (N=104, 51% Female)</u>							
Mean	84.23	4.09	2.08	2.67	43.85	14.53	6.22
SD	2.89	1.82	1.64	1.57	20.51	11.92	2.96
Skewness	.65	.04	.75	.59	-.19	.44	.74
<u>Total (N=95,199, 60% Female)</u>							
Mean	33.55	4.97	4.09	4.90	55.87	26.15	8.57
SD	12.48	1.48	1.91	1.64	18.50	9.79	3.36
Skewness	.96	-.21	-.35	.04	-.35	-1.03	-.01

	All Parameters	Fix Factor	Also Fix	Fix Intercepts	Fix Intercepts
<u>Age Groups</u>	<u>Free</u>	<u>Loadings Equal</u>	<u>Residuals Equal</u>	<u>But Not Residuals</u>	<u>and Residuals</u>
Chi-Squared (df)	1,456.81 (70)	1,648.81 (122)	3,013.11 (187)	8,748.19 (174)	10,383.93 (239)
Log Likelihood (# par)	-638,943.70 (210)	-639,044.90 (158)	-639,746.26 (93)	-642,709.21 (106)	-643,427.54 (41)
AIC	1,278,307.39	1,278,405.80	1,279,678.50	1,285,630.42	1,286,937.09
Sample Size-Adjusted BIC	1,279,627.39	1,279,398.94	1,280,263.07	1,286,296.70	1,287,194.80
Comparative Fit Index	.967	.964	.797	.797	.759
Tucker-Lewis Index	.934	.958	.836	.836	.859
RMSEA	.054	.043	.085	.085	.079
		OK	No	No	No
<u>Level of Education</u>					
Chi-Squared (df)	1,832.85 (35)	1,972.75 (59)	2,115.98 (89)	2,895.94 (83)	3,047.14 (113)
Log Likelihood (# par)	-634,567.12 (105)	-634,617.69 (81)	-634,662.70 (51)	-635,082.62 (57)	-635,129.55 (27)
AIC	1,269,344.23	1,269,397.37	1,269,427.42	1,270,279.24	1,270,313.10
Sample Size-Adjusted BIC	1,270,335.85	1,269,904.92	1,269,746.97	1,270,636.40	1,270,482.28
Comparative Fit Index	.962	.960	.957	.957	.938
Tucker-Lewis Index	.925	.952	.967	.967	.962
RMSEA	.062	.049	.041	.041	.044
		OK	OK	No	No
<u>Sex</u>					
Chi-Squared (df)	1,876.78 (10)	1,917.63 (14)	2,019.93 (19)	3,159.32 (18)	3,278.90 (23)
Log Likelihood (# par)	-629,927.82 (30)	-629,944.21 (26)	-629,972.75 (21)	-630,564.46 (22)	-630,597.05 (17)
AIC	1,259,915.64	1,259,940.42	1,259,987.51	1,261,172.92	1,261,228.09
Sample Size-Adjusted BIC	1,260,103.35	1,260,103.10	1,260,118.91	1,261,310.58	1,261,334.46
Comparative Fit Index	.962	.961	.959	.936	.934
Tucker-Lewis Index	.924	.945	.957	.929	.943
RMSEA	.064	.054	.048	.061	.055
		OK	Not Really	No	No
Note: RMSEA is Root Mean Square Error of Approximation.					

Table 3
Residual Variances and Intercepts of Variables from Age Group Memory Models

<u>Age Group</u>	<u>N</u>	<u>Feature Binding</u>	<u>Visual Pattern Span</u>	<u>Digit Span</u>	<u>Working Memory Span</u>	<u>Spatial Orientation</u>
<u>Residual Variances</u>						
1. 20 & under	11,595	.591	.691	.712	.680	.922
2. 21-25	19,228	.664	.689	.739	.704	.925
3. 26-30	17,146	.649	.680	.720	.715	.857
4. 31-35	12,837	.645	.635	.715	.722	.806
5. 36-40	9,772	.626	.628	.728	.728	.759
6. 41-45	7,426	.661	.601	.732	.783	.686
7. 46-50	6,094	.669	.601	.745	.794	.647
8. 51-55	4,737	.623	.546	.768	.796	.615
9. 56-60	3,166	.647	.522	.766	.789	.575
10. 61-65	1,492	.604	.534	.815	.796	.551
11. 66-70	947	.528	.438	.851	.719	.545
12. 71-75	447	.565	.505	.810	.690	.508
13. 76-80	210	.434	.650	.658	.666	.482
14. 81 & over	104	.443	.644	1.006	.678	.622
<u>Intercepts</u>						
1. 20 & under	11,595	.262	.383	-.104	.137	.021
2. 21-25	19,228	.207	.312	.030	.116	.072
3. 26-30	17,146	.132	.181	.062	.078	.104
4. 31-35	12,837	.064	.061	.079	.027	.089
5. 36-40	9,772	-.028	-.092	.036	-.016	.032
6. 41-45	7,426	-.169	-.273	.008	-.080	-.071
7. 46-50	6,094	-.300	-.460	-.029	-.162	-.143
8. 51-55	4,737	-.441	-.603	-.072	-.229	-.184
9. 56-60	3,166	-.561	-.748	-.091	-.267	-.282
10. 61-65	1,492	-.744	-.905	-.186	-.402	-.414
11. 66-70	947	-.981	-1.134	-.386	-.630	-.610
12. 71-75	447	-1.121	-1.213	-.532	-.832	-.671
13. 76-80	210	-1.105	-1.309	-.482	-.827	-.724
14. 81 & over	104	-1.119	-1.357	-.650	-1.007	-.699
Total	95,201					

Note: Test scores were standardized in the full sample. Residual variances and intercepts presented here were not further standardized within groups.

Table 4
Residual Variances and Intercepts of Variables from Education and Sex Group Memory Models

<u>Education Groups</u>	<u>N</u>	<u>Feature Binding</u>	<u>Visual Pattern Span</u>	<u>Digit Span</u>	<u>Working Memory Span</u>	<u>Spatial Orientation</u>
				<u>Intercepts</u>		
None	488	-.73	-.65	-.72	-.76	-.53
Primary	563	-.46	-.38	-.48	-.51	-.29
Secondary	22,446	-.05	-.07	-.10	-.12	-.06
Technical College	12,168	-.11	-.13	-.14	-.14	-.07
Other College	16,168	.03	.02	-.07	.00	-.02
Graduate First Degree	25,035	.09	.12	.12	.12	.09
Postgraduate Degree	16,474	.00	-.01	.15	.11	.05
Total	93,342					
<u>Sex</u>				<u>Residuals</u>		
Male	37,246	.66	.69	.78	.73	.81
Female	55,273	.61	.66	.67	.73	.80
				<u>Intercepts</u>		
Male	37,246	.02	.13	.10	-.01	.08
Female	55,273	-.01	-.08	-.07	.01	-.06
Total	92,519					

Note: Test scores were standardized in the full sample. Intercepts presented here were not further standardized within groups. Reported education levels from people who reported ages too young to have earned them were considered missing. Many participants did not report sex.

Table 5
Proportions of Variance Attributable to Memory Factor by Age

<u>Age</u>	<u>Feature Binding</u>	<u>Visual Pattern Span</u>	<u>Digit Span</u>	<u>Working Memory Span</u>	<u>Spatial Orientation</u>
20 & under	.310	.243	.254	.250	.157
21-25	.288	.240	.244	.244	.161
26-30	.303	.246	.259	.251	.180
31-35	.314	.261	.270	.255	.192
36-40	.338	.265	.295	.266	.191
41-45	.300	.253	.270	.248	.195
46-50	.301	.258	.261	.248	.195
51-55	.316	.271	.257	.255	.201
56-60	.311	.269	.256	.263	.218
61-65	.311	.292	.247	.233	.206
66-70	.339	.339	.266	.275	.208
71-75	.251	.274	.208	.243	.201
76-80	.446	.293	.261	.331	.194
81 & over	.395	.285	.174	.315	.192

Caption for Figure 1 –

Figure 1 – Normalized test means and standard deviations by test and age group. The lines centered around 0 are means; those centered around 1 are standard deviations.

Caption for Figure 2 –

Figure 2 – Residual variances by test and age group, from fitted regression lines.

Caption for Figure 3 –

Figure 3 – Intercepts by test and age group, from fitted regression lines.

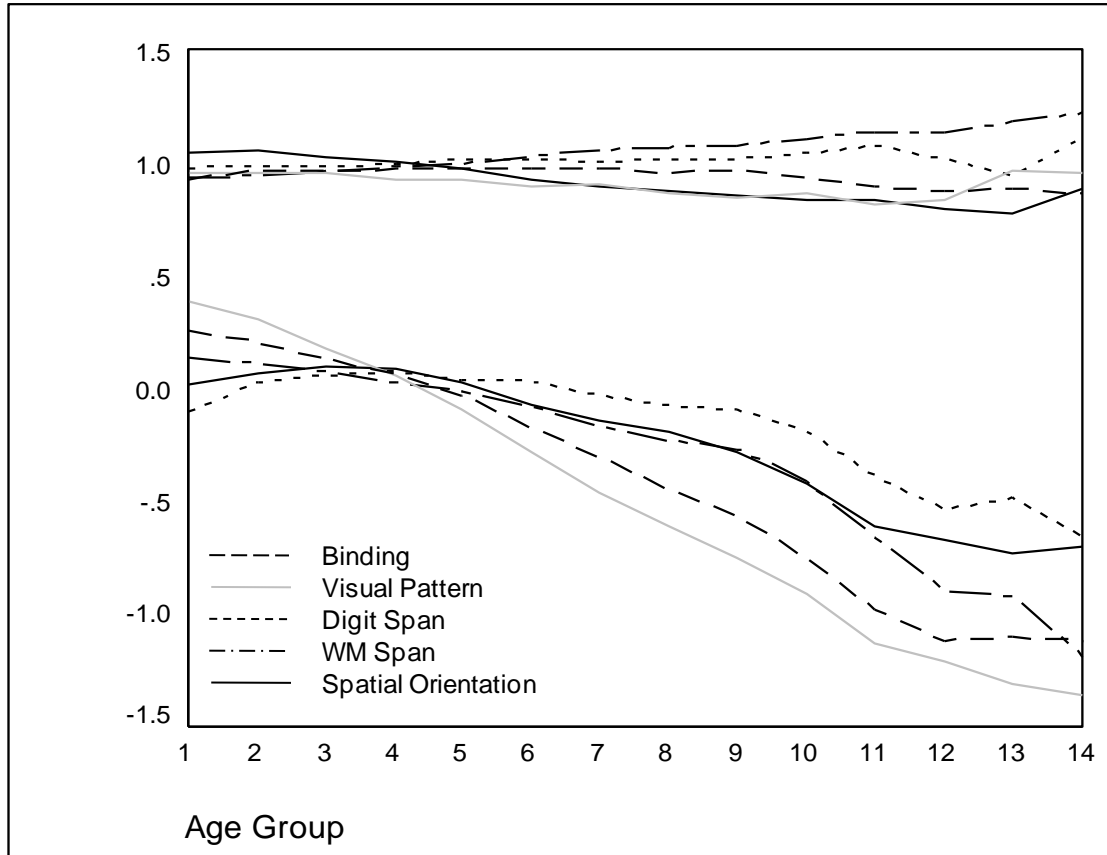


Figure 1

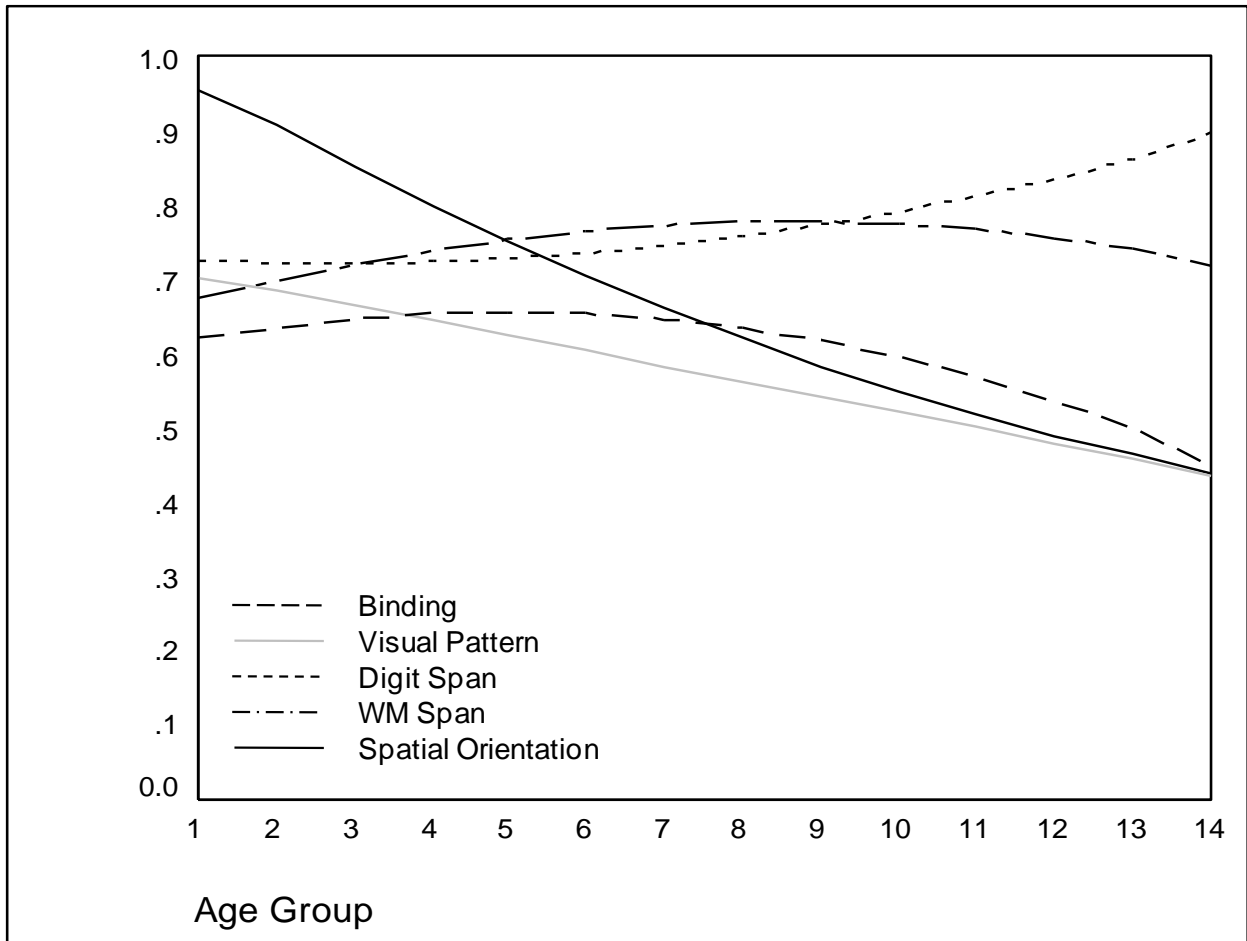


Figure 2

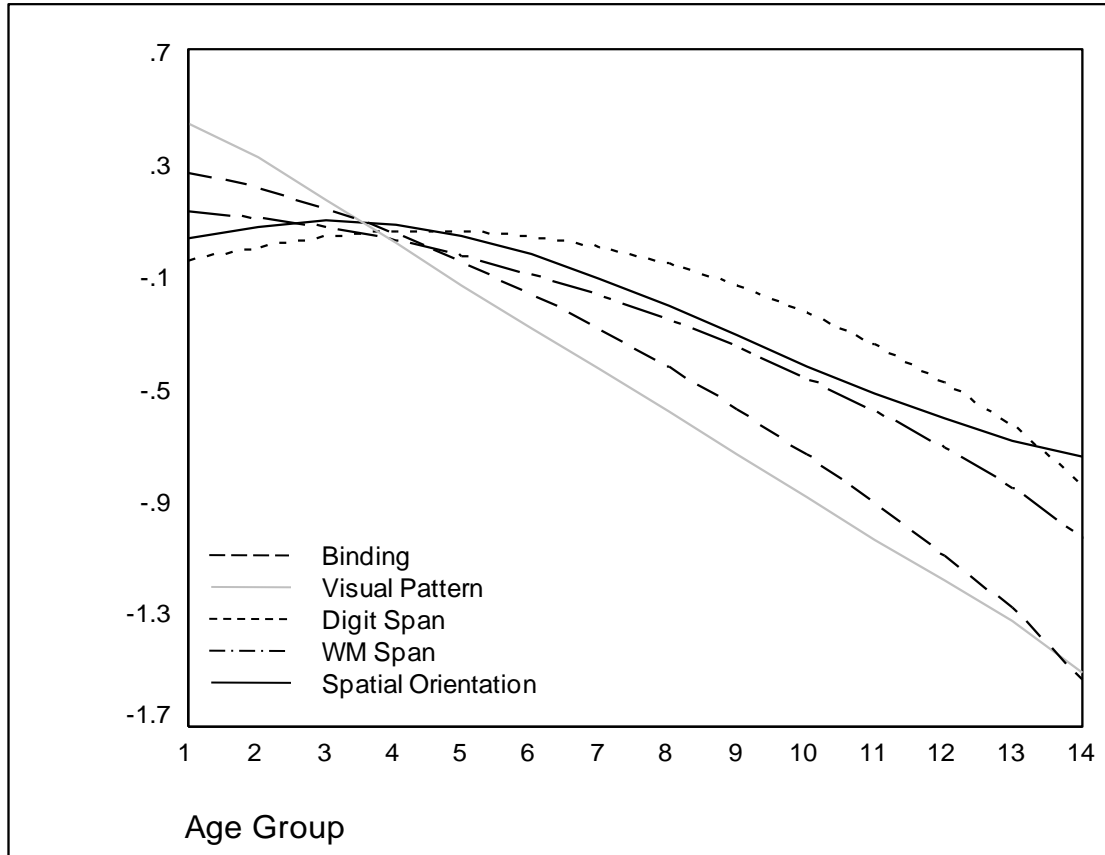


Figure 3